

Today:

- Derandomization via method of Conditional Expectations
- Random Walks!

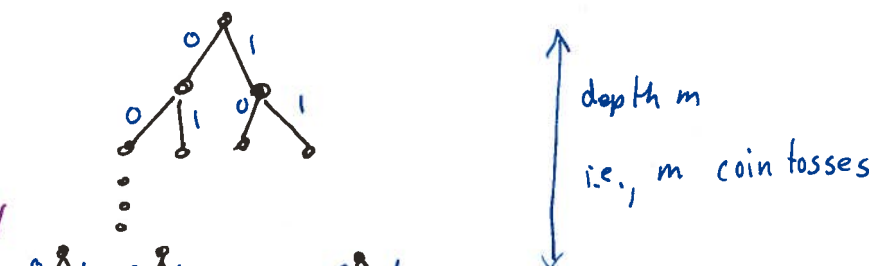
Markov chains + random walks on graphs

Stationary Distributions

More derandomization: The method of conditional expectations

idea: view coin tosses of algorithm as path down a tree of depth m ← # coin tosses

$0 = H$
 $1 = T$



good = correct/reach without good approx pass...

Alternatively: cut values

good randomized algorithm \Leftrightarrow most leaves are good

idea find a good path to leaf "bit-by-bit"

more formally:

Fix randomized algorithm A
input x
 $m = \#$ random bits used by A on x

for $1 \leq i \leq m$ & $r_1, \dots, r_i \in \{0, 1\}$, let $p(r_1, \dots, r_i) =$ fraction of continuations that end in "good" leaf

$C(r_1, \dots, r_i) =$ ave cut value if set i nodes to r_1, \dots, r_i

$$p(r_1, \dots, r_i) = \frac{1}{2} \cdot p(r_1, \dots, r_i, 0) + \frac{1}{2} \cdot p(r_1, \dots, r_i, 1)$$

$$= \Pr_{R_{i+1} \dots R_m} [A(x; r_1, \dots, r_i, R_{i+1}, \dots, R_m) \text{ correct}]$$

$$= \frac{1}{2} \left[\Pr_{R_{i+2} \dots R_m} [A(x; r_1, \dots, r_i, 0, R_{i+2}, \dots, R_m) \text{ correct}] + \Pr_{R_{i+2} \dots R_m} [A(x; r_1, \dots, r_i, 1, R_{i+2}, \dots, R_m) \text{ correct}] \right]$$

by averaging, \exists setting of r_{i+1} to 0 or 1

st. $p(r_1, \dots, r_{i+1}) \geq p(r_1, \dots, r_i)$ Can we figure this out?

if $p(r_1, \dots, r_{i+1}) \geq p(r_1, \dots, r_i) \quad \forall i$

then $p(r_1, \dots, r_m) \geq p(r_1, \dots, r_{m-1}) \geq \dots \geq p(r_1) \geq p(\perp) (\geq 2/3?)$

↑
this is a leaf
so value is 1 or 0,
but if $\geq 2/3$
must be 1

↑
fraction
of good paths

main issue: how do we figure out the best setting of r_{i+1} at each step?

An example: Max Cut (another way to derandomize)

recall algorithm:

flip n coins r_1, \dots, r_n
put node i in S if $r_i = 0$ & T if $r_i = 1$
output S, T

derandomization:

expected
size of
cut

$$e(r_1, \dots, r_i) = E_{R_{i+1}, \dots, R_N} [| \text{cut}(S, T) | \text{ given } r_1, \dots, r_i \text{ choices made}]$$

$$e(\perp) \geq \frac{|E|}{2} \quad (\text{from previous lecture})$$

↑
no choices fixed

how do we calculate $e(r_1, \dots, r_{i+1})$?

Let

$$S_{i+1} = \left\{ \begin{array}{l} \text{nodes} \\ j \mid j \leq i+1, r_j = 0 \end{array} \right\}$$

$$T_{i+1} = \left\{ \begin{array}{l} \text{nodes} \\ j \mid j \leq i+1, r_j = 1 \end{array} \right\}$$

$$U_{i+1} = \left\{ \begin{array}{l} \text{nodes} \\ j \mid j \geq i+2 \text{ and } j \leq n \end{array} \right\}$$

} S + T so far

"Undecided"

So

fact: $e(r_1, \dots, r_{i+1}) = (\# \text{ edges between } S_{i+1} + T_{i+1}) + \frac{1}{2} (\# \text{ edges touching } U_{i+1})$
 (follows from same argument as last lecture)

Note: don't need to calculate $e(r_1, \dots, r_{i+1})$

just need to compare $e(r_1, \dots, r_i, 0)$ vs. $e(r_1, \dots, r_i, 1)$ - is it $\{<, >\}$

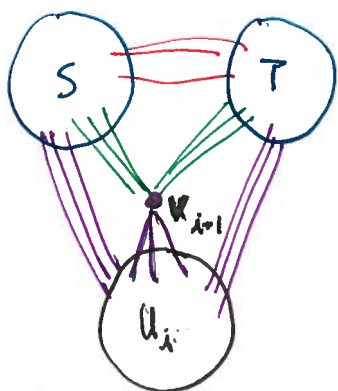
- note:
- # edges between $S_{i+1} + T_{i+1}$ same for both
 - U_i term is same for both

- first term differs only on edges adjacent to node $i+1$



to maximize this, place node $i+1$ to maximize cut size

i.e. $|\# \text{ edges between node } i+1 + S_i|$
 vs. $|\# \text{ edges between node } i+1 + T_i|$



Corresponds to :

Greedy Algorithm :

1) $S \leftarrow \emptyset, T \leftarrow \emptyset$

2) For $i=0 \dots N-1$

place node i in S if $\# \text{edges between } i \text{ and } T$
 $\geq \# \text{edges between } i \text{ and } S$

else place in T

Random walks

Markov chains :

Ω = set of "states" (or nodes) (here always FINITE)

$X_0 \dots X_t \in \Omega$ sequence of visited states

Markovian property :

$$\Pr[X_{t+1} = y \mid X_0 = x_0, X_1 = x_1, X_2 = x_2, \dots, X_t = x_t] \\ = \Pr[X_{t+1} = y \mid X_t = x_t]$$

} Next step depends only on where you are. Not how you got there

Wlog, assume transitions independent of time :

i.e. $P(x, y) = \Pr[X_{t+1} = y \mid X_t = x]$

so can use "transition matrix" to represent it

Important special case :

transitions uniform on subset corresponding to neighbors of node

def. random walk on $G = (V, E)$

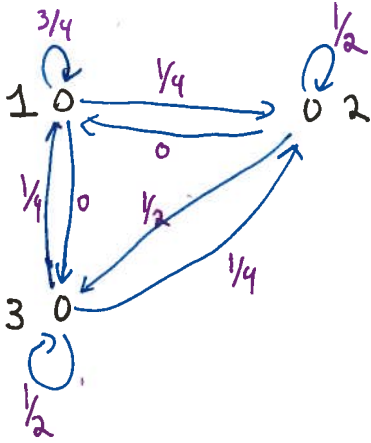
is a sequence s_0, s_1, \dots of nodes

where s_0 is a start node.

At each step i , s_{i+1} picked uniformly from $N(s_i)$
outedges

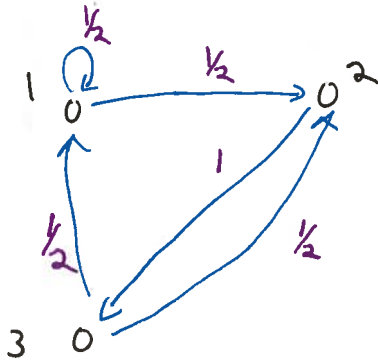
examples

Markov chain



$$p: \begin{array}{c} 1 \\ 2 \\ 3 \end{array} \begin{array}{|c|c|c|} \hline 1 & 2 & 3 \\ \hline 3/4 & 1/4 & 0 \\ \hline 0 & 1/2 & 1/2 \\ \hline 1/4 & 1/4 & 1/2 \\ \hline \end{array}$$

random walk on digraph



$$p: \begin{array}{c} 1 \\ 2 \\ 3 \end{array} \begin{array}{|c|c|c|} \hline 1 & 2 & 3 \\ \hline 1/2 & 1/2 & 0 \\ \hline 0 & 0 & 1 \\ \hline 1/2 & 1/2 & 0 \\ \hline \end{array}$$
 $d(i) = \# \text{ outedges of node } i$

$$P(i,j) = \begin{cases} \frac{1}{d(i)} & \text{if } (i,j) \in E \\ 0 & \text{o.w.} \end{cases}$$

$$\forall i \quad \sum_j P(i,j) = 1$$

Distributions after t steps

Transition probabilities for t steps: $P^t(x,y) = \begin{cases} P(x,y) & t=1 \\ \sum_z P(x,z)P^{t-1}(z,y) & t > 1 \end{cases}$ } matrix multiplication. $P^t = \underbrace{P \times P \times \dots \times P}_{t \text{ times}}$

Initial distribution: $\pi^{(0)} = (\pi_1^{(0)}, \dots, \pi_n^{(0)})$ where $\pi_i^{(0)} = P_i$ [start at node i]

distribution after one step:

$$\pi^{(1)} = \pi^{(0)} \cdot P = \left(\sum_z P(z,1) \cdot \pi(z), \sum_z P(z,2) \pi(z), \dots \right)$$

⋮

t-step distribution: $\pi^{(t)} = \pi^{(0)} \cdot P^t$

Finite Markov Chain Properties

Stochastic matrix: rows of P sum to 1

all M.C.'s have this property

doubly stochastic matrix: rows & columns sum to 1

e.g. random walk on undirected ~~d-regular~~ graphs or digraph in which indegree = outdegree = const for all nodes

not even all interesting M.C.'s satisfy this

irreducible: ("strongly connected")

$$\forall x,y \exists t = t(x,y) \text{ st. } P^t(x,y) > 0$$

ergodic: $\exists t_0$ st. $\forall t > t_0 \forall x,y P^t(x,y) > 0$

← stronger than irreducible! why?

Aperiodic:
 $\forall x \quad \text{gcd} \{t : P^t(x,x) > 0\} = 1$

gcd of "possible" cycle length = 1

not bipartite, k-partite...

Thm Ergodic \Leftrightarrow Irreducible + Aperiodic

Stationary Distributions

does it depend on π_0 ?

Stationary distribution π

$\forall y \quad \pi(y) = \sum_x \pi(x) P(x,y)$

so $\pi^{(t)} = \pi^{(t-1)}$

Will consider P s.t. π is unique & exists } i.e. doesn't depend on π_0

if periodic: could have no stat. dist. or several
 if reducible: could have lots of stat. dist.

if $\pi_0 = (0,1)$
 then $\pi_{2i} = (0,1)$
 $\pi_{2i+1} = (1,0)$

Some stat dist's:
 $(\frac{1}{2}, \frac{1}{2}) \quad (0,1) \quad (1,0) \dots$

Important Thm every ergodic M.C. has unique stationary distribut